

Model Based Statistics in Biology.

Part IV. The General Linear Model. Multiple Explanatory Variables.

Chapter 14.2 ANCOVA - Statistical Control

ReCap.	Part I (Chapters 1,2,3,4), Part II (Ch 5, 6, 7)
ReCap	Part III (Ch 9, 10, 11)
ReCap	Multiple Regression (Ch 12)
ReCap	Multiple Categorical Variables (Ch 13)
14.1	Comparing Regression Lines
14.2	Statistical Control
14.3	Model Revision
14.4	More than two explanatory variables (to be written)

CrwTb9_1.xls Ch14.xls

on chalk board

ReCap Part I (Chapters 1,2,3,4) Quantitative reasoning is based on models, including statistical analysis based on models.

ReCap Part II (Chapters 5,6,7)

Hypothesis testing uses the logic of the null hypothesis to declare a decision.

Estimation is concerned with the specific value of an unknown population parameter.

ReCap (Ch 9, 10,11) The General Linear Model with a single explanatory variable.

ReCap (Ch 12) GLM with more than one regression variable (multiple regression)

ReCap (Ch 13) GLM with more than one categorical variable (ANOVA).

ReCap (Ch 14) ANCOVA with GLM - Comparing regression lines.

Today: Statistical control, with ANCOVA.
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Statistical control allows the effects of one variable to be removed, in order to arrive at a better analysis of the effects of another variable.

Wrap-up.

Statistical control improves analysis by removing the effects of a secondary variable, to achieve lower residual mean square and better analysis of the variable of interest.

In ANCOVA either the ratio scale or the nominal scale explanatory variable can be the control variable. A ratio scale response variable (*e.g.* fish production from lakes) can be analyzed relative to a ratio scale explanatory variable (*e.g.* size of lake) controlled for a nominal scale variable (*e.g.* temperate versus tropical lakes). Or a nominal scale explanatory variable (*e.g.* experimental treatment versus control) can be tested controlling for the effects of a ratio scale explanatory variable (*e.g.* metabolic rate of the animal).

Of these two possibilities, the more commonly encountered is that of a classification (nominal scale) explanatory variable, controlled for a ratio scale variable. An example of this was worked through today.

Introduction.

ANCOVA is applied to data situations that have a mixture of both ratio and nominal scale explanatory variables. We have already looked at ANCOVA where we compare slopes of one or more regression lines, using the interaction term in the ANCOVA model. Today we will look at another application of ANCOVA, where we compare several groups (ANOVA explanatory variable) controlling for the effects of a second explanatory variable (regression variable on a ratio type of scale.). To do this analysis we will need to establish that the slopes are the same in the groups (no interaction term).

Data from Table 9.1 in M.J. Crawley (1993) *GLIM for Ecologists*.

The data consist of seed production in 40 plants allocated at random to two treatments, grazed or not grazed by rabbits.

The grazed plants were exposed to rabbits during the first two weeks of stem elongation, then protected from subsequent grazing.

The size of the plant was thought to influence seed production so the diameter at the top of the root stock (in mm) was measured before exposure to grazing.

At end of growing season, fruit production ($M_{\text{fruit}} = \text{mg dry wt}$) was recorded for each of the 40 plants.

1. Construct model

Verbal model.

Fruit production depends on grazing and root size.
Is there evidence for a difference in fruit production between grazed and ungrazed plants, if we control for the relation to root size?

fruit (mg)	root (mm)	grazed
59.77	6.225	n
60.98	6.487	n
14.73	4.919	n
19.28	5.13	n
34.25	5.417	n
35.53	5.359	n
87.73	7.614	n
63.21	6.352	n
24.25	4.975	n
64.34	6.93	n
52.92	6.248	n
32.35	5.451	n
53.61	6.013	n
54.86	5.928	n
64.81	6.264	n
73.24	7.181	n
80.64	7.001	n
18.89	4.426	n
75.49	7.302	n
46.73	5.836	n
80.31	8.988	y
82.35	8.975	y
105.1	9.844	y
73.79	8.508	y
50.08	7.354	y
78.28	8.643	y
41.48	7.916	y
98.47	9.351	y
40.15	7.066	y
116.1	10.25	y
38.94	6.958	y
60.77	8.001	y
84.37	9.039	y
70.11	8.91	y
14.95	6.106	y
70.7	7.691	y
71.01	8.515	y
83.03	8.53	y
52.26	8.158	y
46.64	7.382	y

2. Execute analysis.

Code the model statement in statistical package according to the GLM

$$M_{seedt} = \beta_o + \beta_{root}Root + \beta_{Gr}Gr + \beta_{root \cdot Gr}Root \cdot Gr + \varepsilon$$

```
MTB > glm 'Mfruit' = 'root' 'Gr' 'root'*'Gr';
SUBC> covariate 'root';
SUBC> fits c4;
SUBC> residuals c5.
```

Fits and residuals from:

model statement output of fitted values and residuals (as above)
 or parameters reported by GLM routine
 or direct calculation of parameters

Here are the parameter estimates.

The overall mean fruit production is $\hat{\beta}_o = 59.41$ mg

The mean for grazed and ungrazed is expressed as a deviation from $\hat{\beta}_o$

$$\hat{\beta}_o + \hat{\beta}_{GR} = \begin{cases} \text{mean}(M_{GR=no}) = 59.41 - 8.53 = 50.88 \\ \text{mean}(H_{GR=yes}) = 59.41 + 8.53 = 67.94 \text{ mg} \end{cases}$$

The slope parameter for grazed and ungrazed together is $\hat{\beta}_{root} = 23.625$ mg/mm

Note that the ANCOVA estimate of the slope differs from the slope estimate by simple regression, without the grazing term in the model.

```
MTB > regress 'fruit' 1 'root'.

The regression equation is
fruit = - 41.3 + 14.0 root

Predictor      Coef      Stdev    t-ratio      p
Constant      -41.31     10.73     -3.85      0.000
root           14.026     1.464      9.58      0.000
```

This is because the ungrazed plants are smaller, hence to the left of the grazed plants in the graph. This lateral offset reduces the overall slope from around 23 mg/mm in each group to 14.0 mg/mm across all the data.

The deviation from the ANCOVA estimate of the overall slope are small.

$$\hat{\beta}_{Root*Gr} = \begin{matrix} -0.371 \text{ mg/mm} \\ +0.371 \text{ mg/mm} \end{matrix}$$

$$\hat{\beta}_{root} + \hat{\beta}_{root*GR} = \begin{cases} \text{Slope}(H_{pers}) = 23.625 - 0.371 = 23.996 \\ \text{Slope}(H_{pseu}) = 23.625 + 0.371 = 23.254 \end{cases}$$

These deviations are symmetrical because there are only two groups.

2. Execute analysis.

Compare to regression equation (one slope and one intercept) for each species:

$$H_{Gr=No} = -94.367 + 23.996 \text{ Root}$$

$$M_{Gr=Yes} = -125.28 + 23.254 \text{ Root}$$

The GLM routine computes fitted and residual values.

3. Evaluate the model

Plot residuals versus fits.

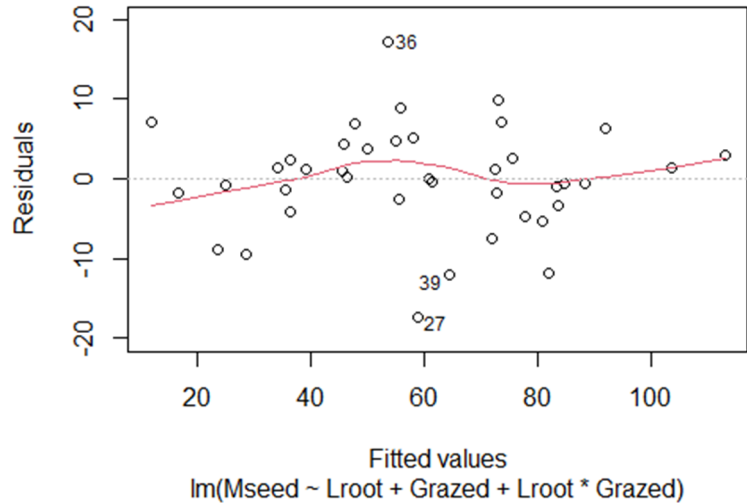
Straight line assumption acceptable. No bowls or arches in plot.

Error model.

If n small, evaluate assumptions for p-values from chisquare (t, F) distributions.

n = 40,

So even substantial deviations will have little distorting effect on calculation of parameter estimates and p-values.



a. Homogeneous? Yes

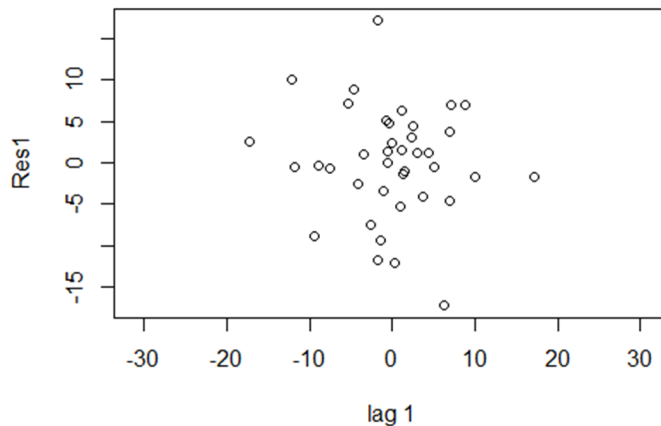
Residuals do not change in any systematic way with fitted values (no cones).

b. Normal?

The residuals look normal plotted as a histogram and in QQ plot

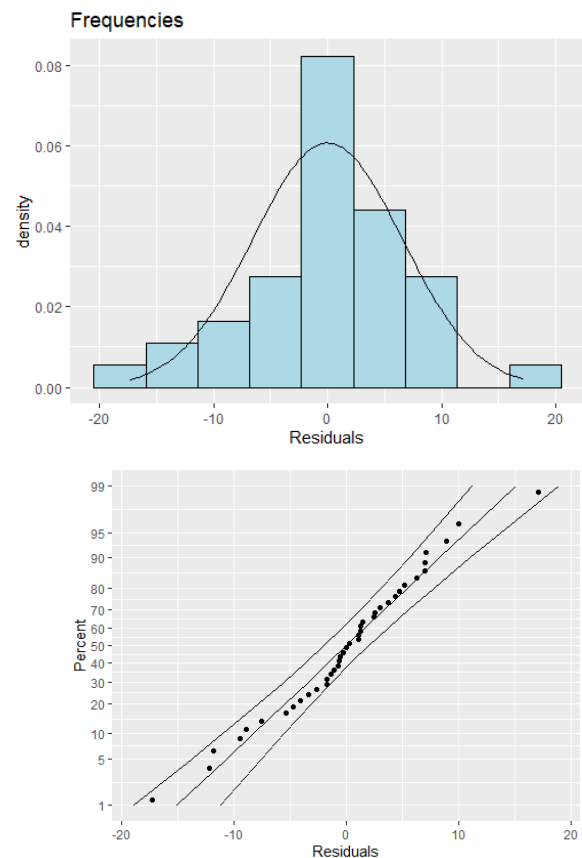
c. Independent?

Each residual plotted against its neighbor.



No evidence of non-independence.

d. Sum(res) = 0? Yes



4. State population and whether sample is representative.

Population might be that from which the plants were selected.

In this example, the population will be taken as all possible measurements, given the protocol.

5. Decide on mode of inference. Is hypothesis testing appropriate?

It is clear that seed production depends on root size. Hypothesis testing is irrelevant. It is not clear whether fruit production depends on grazing, after controlling for effects of root size. Hypothesis testing appropriate.

6. State H_A H_0 pairs, test statistic, distribution, tolerance for Type I error.

Terms in model.

We begin with the interaction term. Are slopes parallel ?

$$H_A: \text{var}(\beta_{Root*Gr}) > 0$$

$$H_0: \text{var}(\beta_{Root*Gr}) = 0$$

This is equivalent to following hypotheses concerning parameters

$$\beta_{root*Gr=0} \neq \beta_{root*Gr=1} \quad (\text{slope not parallel})$$

$$\beta_{root*Gr=0} = \beta_{root*Gr=1} \quad (\text{slopes parallel})$$

7. ANOVA

```

MTB > glm 'seed' = 'root' 'grazing' 'root'*'grazing';
SUBC> covariate 'root';
SUBC> fits c8;
SUBC> residuals c9.

Factor   Levels Values
grazing    2      0      1

Analysis of Variance for seed

Source      DF      Seq SS      Adj SS      Adj MS      F      P
root         1      16800.4      18791.6      18791.6      402.57  0.000
grazing      1       5266.7       157.1       157.1       3.37  0.075
grazing*root 1         4.6         4.6         4.6       0.10  0.754
Error       36      1680.5      1680.5       46.7
Total       39      23752.2

```

While we can reject the null hypothesis, we cannot make a statement about the alternative hypothesis, concerning the interaction term. We look at the measure of evidence.

$$LR = (4.6/1680.5)^{-40/2} \ll 1.$$

There is no evidence of an interactive effect. The slopes are parallel.

8. When assumptions not met, decide whether to re-compute likelihood ratio.

Assumptions were met, so continue to next step.

9. Conclusion The slopes are parallel. $\beta_{root*Gr=0} = \beta_{root*Gr=1}$

There was no evidence for interactive effect, so we could examine the grazing term.

It is close to the 5% criterion. Note, however the substantial difference between the Seq SS and the Adj SS of the grazing term.

Because there is no evidence for an interactive effect, we remove the interaction term to increase the power of the analysis. Revise the model. Back to step 1.

1. Construct Model

$$Mseed = \beta_o + \beta_{root} * Root + \beta_{Gr} * Gr + \varepsilon$$

This is our model to test for grazing effects controlled for plant size (root diameter). The interaction term has been removed.

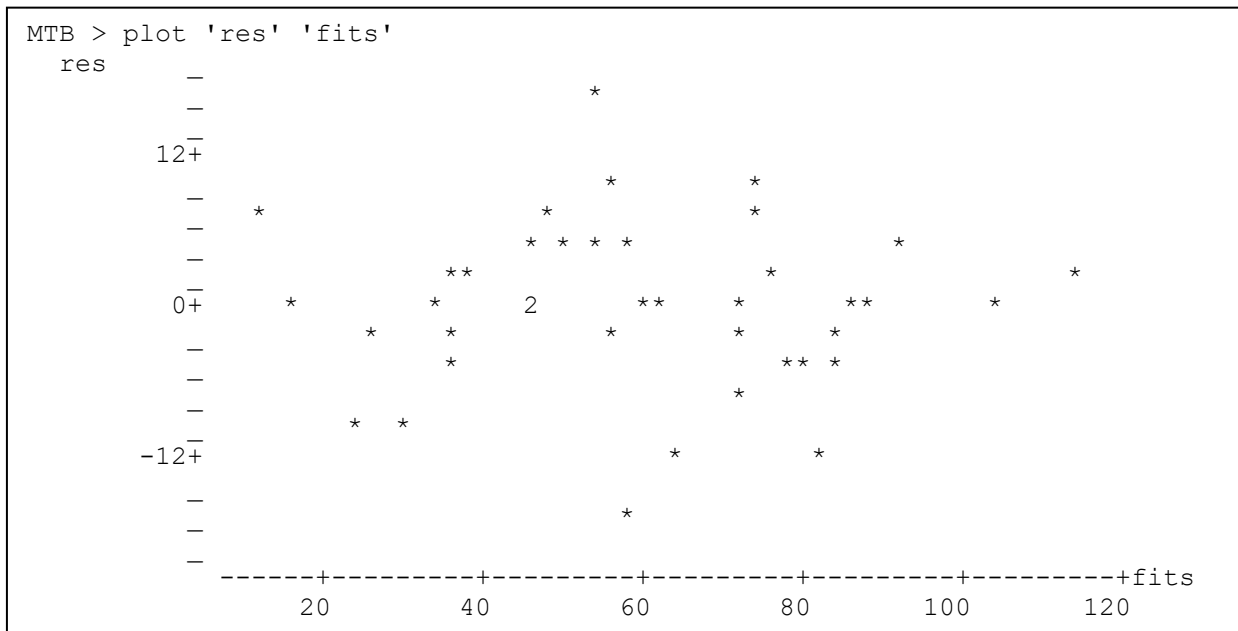
ANCOVA' routines aimed at statistical control assume no interaction term. The

In this example, we tested the assumption of no interaction, rather than blindly assuming it to be true.

2. Execute analysis.

3. Evaluate model

Plot residuals vs fits



1. Straight line acceptable, no bowls or arches.
2. a. Var(error) = constant ? Yes. No cones.
b. Normal errors ? Yes. Histogram OK, so no further diagnosis
c. Independent errors ? Yes (not shown)

4. Partition df and SS according to model.

No change

$$LR = (5267/1685)^{-40/2} \gg 1000$$

Strong sequential and with adjusted sum of squares now the same.

Strong evidence for grazing effect, when root size included in analysis.

5. Hypothesis testing? Yes.

6. State hypothesis H_A / H_0

Terms in model. Only one term will be examined, the grazing effect.

$$H_A: \text{Var}(\beta_{Gr}) > 0$$

$$H_0: \text{Var}(\beta_{Gr}) = 0$$

Equivalent to following hypotheses for parameters.

$$H_A: \beta_{Gr=0} \neq \beta_{Gr=1} \text{ (grazing affect growth, controlled for size)}$$

$$H_0: \beta_{Gr=0} = \beta_{Gr=1}$$

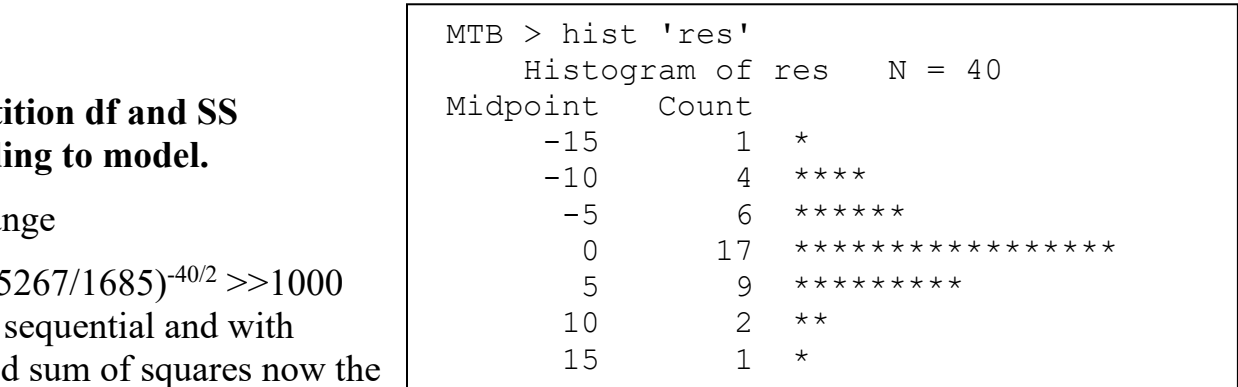
We can state a more specific hypothesis about the parameter, based on the biology.

$$H_A: \beta_{Gr=0} > \beta_{Gr=1} \text{ (grazing reduces growth, controlled for size)}$$

$$H_0: \beta_{Gr=0} \leq \beta_{Gr=1}$$

We are not interested in testing whether seed production depends on root size, it is obvious from the plot that it does.

7. ANOVA Table.



```

MTB > glm 'seed' = 'root' 'grazing';
SUBC> covariate 'root';
SUBC> fits c8;
SUBC> residuals c9.
Factor  Levels Values
grazing  2      0      1
Analysis of Variance for seed

Source      DF      Seq SS      Adj SS      Adj MS      F      P
root        1      16800      19155      19155      420.60  0.000
grazing     1       5267       5267       5267      115.64  0.000
Error      37       1685       1685         46
Total      39      23752
    
```

8. Recompute Type I error?

No. Assumptions met.

9. Declare decision.

Reject H_0 . The observed difference in growth, controlled for root size, is not due to chance. $F_{1,37} = 115.64$ $p < 0.00001$

10. Analysis of parameters of biological interest.

	grazing	N	MEAN	MEDIAN	TRMEAN	STDEV	SEMEAN
fruit	0	20	50.88	54.24	50.84	21.76	4.87
	1	20	67.94	70.85	68.21	24.97	5.58

When root size is not taken into account, the fruit production appears to be less for ungrazed than for grazed.

Ungrazed 50.88 mg
 Grazed -67.94 mg
 Difference -17.06 mg

This is because the grazed plants were larger than the ungrazed plants.

To compare grazed vs ungrazed, controlled for size, we calculate the vertical separation between the two regression lines. The most convenient point at which to do this is the point at which $x = \text{zero}$ (the y -intercepts).

$\hat{\alpha}$	=	β_o	-	β_{root}	*	mean(X)
$\hat{\alpha}_{Gr=no}$	=	Mean($M_{Gr=no}$)	-	β_{root}	*	Mean($root_{Gr=no}$)
	=	50.88	-	23.6	*	6.053
	=	-91.729 mg				
$\hat{\alpha}_{Gr=yes}$	=	Mean($M_{Gr=Yes}$)	-	β_{root}	*	Mean($root_{Gr=Yes}$)
	=	67.94	-	23.6	*	8.309
	=	-127.82 mg				

The intercept for grazed is below that for ungrazed.

The vertical separation between the two regression lines is:

Ungrazed -91.729 mg
 Grazed -(-127.820) mg
 Difference 36.091 mg

When root size is taken into account, the fruit production for grazed plants is less than for ungrazed. The fruit production for grazed plants was less by 36 mg.